## Statistical Analysis

Lecture 02

## Books

## PowerPoint

http://www.bu.edu.eg/staff/ahmedaboalatah14-courses/14767


## Agenda

>Review "The Central Limit Theorem"
>Sampling Distribution of the Difference between Two Means
$>$ Sampling Distribution of $S^{2}$

# Sampling Distributions and Data Descriptions 

CHAPTER 8

## The Central Limit Theorem

Central Limit Theorem: If $\bar{X}$ is the mean of a random sample of size $n$ taken from a population with mean $\mu$ and finite variance $\sigma^{2}$, then the limiting form of the distribution of

$$
Z=\frac{\bar{X}-\mu}{\sigma / \sqrt{n}}
$$

as $n \rightarrow \infty$, is the standard normal distribution $n(z ; 0,1)$.

## Ex 1:

1. Suppose the grades in a mathematics class are Normally distributed with a mean of 75 and a standard deviation of 6.

What is the probability that the average grade for 9 randomly selected students was at least 80?

## Solution :

In this case, $\mu=75$ and $\sigma=6$.
We need to calculate the probability $\mathrm{P}(\bar{X}>80)$ with $\mathrm{n}=9$.

$$
\begin{gathered}
P(\bar{X} \geq 80)=P\left(\frac{\bar{X}-\mu}{\frac{\sigma}{\sqrt{n}}} \geq \frac{80-75}{\frac{6}{\sqrt{9}}}\right)=P(Z \geq 2.5) \\
P(Z \geq 2.5)=1-P(Z<2.5)=P(Z \leq-2.5) \\
=0.0062
\end{gathered}
$$

# Sampling Distribution of the Difference between Two Means 

## Sampling Distribution of the Difference between Two Means

Suppose that we have two populations:
The first with mean $\mu_{1}$ and variance $\sigma^{2}{ }_{1}$, and the second with mean $\mu_{2}$ and variance $\sigma^{2}{ }_{2}$.

Let the statistic $\bar{X}_{1}$ represent the mean of a random sample of size $\mathrm{n}_{1}$ selected from the first population, and

The statistic $\bar{X}_{2}$ represent the mean of a random sample of size $n_{2}$ selected from the second population, independent of the sample from the first population.

## Sampling Distribution of the Difference between Two Means

According to Theorem 8.2, the variables $X_{1}$ and $X_{2}$ are both approximately normally distributed with means $\mu_{1}$ and $\mu_{2}$ and variances $\sigma_{1}^{2} / n_{1}$ and $\sigma_{2}^{2} / n_{2}$, respectively. This approximation improves as $n_{1}$ and $n_{2}$ increase.

## Sampling Distribution of the Difference between Two Means

By choosing independent samples from the two populations we ensure that the variables $\bar{X}_{1}$ and $\bar{X}_{2}$ will be independent, and then using Theorem 7.11 , with $a_{1}=1$ and $a_{2}=-1$, we can conclude that $\bar{X}_{1}-\bar{X}_{2}$ is approximately normally distributed with mean

$$
\mu_{\bar{X}_{1}-\bar{X}_{2}}=\mu_{\bar{X}_{1}}-\mu_{\bar{X}_{2}}=\mu_{1}-\mu_{2}
$$

and variance

$$
\sigma_{\bar{X}_{1}-\bar{X}_{2}}^{2}=\sigma_{\bar{X}_{1}}^{2}+\sigma_{\bar{X}_{2}}^{2}=\frac{\sigma_{1}^{2}}{n_{1}}+\frac{\sigma_{2}^{2}}{n_{2}}
$$

## Theorem 8.3:

If independent samples of size $n_{1}$ and $n_{2}$ are drawn at random from two populations, discrete or continuous, with means $\mu_{1}$ and $\mu_{2}$ and variances $\sigma_{1}^{2}$ and $\sigma_{2}^{2}$, respectively, then the sampling distribution of the differences of means, $\bar{X}_{1}-\bar{X}_{2}$, is approximately normally distributed with mean and variance given by

$$
\mu_{\bar{X}_{1}-\bar{X}_{2}}=\mu_{1}-\mu_{2} \text { and } \sigma_{\bar{X}_{1}-\bar{X}_{2}}^{2}=\frac{\sigma_{1}^{2}}{n_{1}}+\frac{\sigma_{2}^{2}}{n_{2}} .
$$

Hence,

$$
Z=\frac{\left(\bar{X}_{1}-\bar{X}_{2}\right)-\left(\mu_{1}-\mu_{2}\right)}{\sqrt{\left(\sigma_{1}^{2} / n_{1}\right)+\left(\sigma_{2}^{2} / n_{2}\right)}}
$$

is approximately a standard normal variable.

## Case Study 8.2: Paint Drying Time:

Paint Drying Time: Two independent experiments are run in which two different types of paint are compared. Eighteen specimens are painted using type $A$, and the drying time, in hours, is recorded for each. The same is done with type $B$. The population standard deviations are both known to be 1.0.

Assuming that the mean drying time is equal for the two types of paint, find $P\left(\bar{X}_{A}-\bar{X}_{B}>1.0\right)$, where $\bar{X}_{A}$ and $\bar{X}_{B}$ are average drying times for samples of size $n_{A}=n_{B}=18$.

Solution: From the sampling distribution of $\bar{X}_{A}-\bar{X}_{B}$, we know that the distribution is approximately normal with mean

$$
\mu_{\bar{X}_{A}-\bar{X}_{B}}=\mu_{A}-\mu_{B}=0
$$

and variance

$$
\sigma_{X_{A}-\bar{X}_{B}}^{2}=\frac{\sigma_{A}^{2}}{n_{A}}+\frac{\sigma_{B}^{2}}{n_{B}}=\frac{1}{18}+\frac{1}{18}=\frac{1}{9} .
$$



Figure 8.5: Area for Case Study 8.2.
The desired probability is given by the shaded region in Figure 8.5. Corresponding to the value $\bar{X}_{A}-\bar{X}_{B}=1.0$, we have

$$
z=\frac{1-\left(\mu_{A}-\mu_{B}\right)}{\sqrt{1 / 9}}=\frac{1-0}{\sqrt{1 / 9}}=3.0
$$

$$
P(Z>3.0)=1-P(Z<3.0)=1-0.9987=0.0013
$$

## EKのn@ OBO.

The television picture tubes of manufacturer $A$ have a mean lifetime of 6.5 years and a standard deviation of 0.9 year, while those of manufacturer $B$ have a mean lifetime of 6.0 years and a standard deviation of 0.8 year. What is the probability that a random sample of 36 tubes from manufacturer $A$ will have a mean lifetime that is at least 1 year more than the mean lifetime of a sample of 49 tubes from manufacturer $B$ ?

## Solution:

We are given the following information:

$$
\begin{array}{cc}
\hline \text { Population 1 } & \text { Population 2 } \\
\hline \mu_{1}=6.5 & \mu_{2}=6.0 \\
\sigma_{1}=0.9 & \sigma_{2}=0.8 \\
n_{1}=36 & n_{2}=49 \\
\hline
\end{array}
$$



If we use Theorem 8.3, the sampling distribution of $\bar{X}_{1}-\bar{X}_{2}$ will be approximately normal and will have a mean and standard deviation

$$
\mu_{\bar{X}_{1}-\bar{X}_{2}}=6.5-6.0=0.5 \quad \text { and } \quad \sigma_{\bar{X}_{1}-\bar{X}_{2}}=\sqrt{\frac{0.81}{36}+\frac{0.64}{49}}=0.189 .
$$

The probability that the mean lifetime for 36 tubes from manufacturer $A$ will be at least 1 year longer than the mean lifetime for 49 tubes from manufacturer $B$ is given by the area of the shaded region in Figure 8.6. Corresponding to the value $\bar{x}_{1}-\bar{x}_{2}=1.0$, we find that

$$
z=\frac{1.0-0.5}{0.189}=2.65,
$$

and hence

$$
\begin{aligned}
P\left(\bar{X}_{1}-\bar{X}_{2} \geq 1.0\right) & =P(Z>2.65)=1-P(Z<2.65) \\
& =1-0.9960=0.0040 .
\end{aligned}
$$

## Sampling Distribution of $S^{2}$

## Sampling Distribution of $S^{2}$

If a random sample of size $n$ is drawn from a normal population with mean $\mu$ and variance $\sigma^{2}$, and the sample variance is computed, we obtain a value of the statistic $S^{2}$. We shall proceed to consider the distribution of the statistic $(n-1) S^{2} / \sigma^{2}$.

Start From :

$$
\sum_{i=1}^{n}\left(X_{i}-\mu\right)^{2}
$$

## Sampling Distribution of $S^{2}$

By the addition and subtraction of the sample mean $\bar{X}$, it is easy to see that

$$
\begin{aligned}
\sum_{i=1}^{n}\left(X_{i}-\mu\right)^{2} & =\sum_{i=1}^{n}\left[\left(X_{i}-\bar{X}\right)+(\bar{X}-\mu)\right]^{2} \\
& =\sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}+\sum_{i=1}^{n}(\bar{X}-\mu)^{2}+2(\bar{X}-\mu) \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right) \\
& =\sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}+n(\bar{X}-\mu)^{2}
\end{aligned}
$$

Dividing each term of the equality by $\sigma^{2}$

$$
\text { and substituting }(n-1) S^{2} \text { for } \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2} \text {, }
$$

we obtain

$$
\frac{1}{\sigma^{2}} \sum_{i=1}^{n}\left(X_{i}-\mu\right)^{2}=\frac{(n-1) S^{2}}{\sigma^{2}}+\frac{(\bar{X}-\mu)^{2}}{\sigma^{2} / n} .
$$

## Sampling Distribution of $S^{2}$

$$
\sum_{i=1}^{n} \frac{\left(X_{i}-\mu\right)^{2}}{\sigma^{2}}
$$

$$
\frac{(\bar{X}-\mu)^{2}}{\sigma^{2} / n}
$$

$$
\frac{(n-1) S^{2}}{\sigma^{2}}
$$

The First term is a chi-squared random variable with n degrees of freedom. ( $X_{1}$ to $X_{n}$ and no relation in the equation)

The second term on the right-hand side is $Z^{2}$, which is a chi-squared random variable with 1 degree of freedom. (only sample average)

The third term is a chi-squared random variable with $\mathrm{n}-1$ degree of freedom. ( $X_{1}$ to $X_{n}$ and relation in the equation between them (average))

## Degrees of Freedom as a Measure of Sample Information

Example : suppose we have three numbers and we know their average is equal to zero then we have freedom to choose two of the three numbers
$X_{1}=5, X_{2}=10$, and average $=0$
But for the third one
There is no freedom to choose $X_{3}$
$X_{3}=3^{*}$ average $-X_{1}-X_{2}$

## Degrees of Freedom as a Measure of Sample Information

as indicating that when $\mu$ is not known

considers the distribution of

$$
\sum_{i=1}^{n} \frac{\left(X_{i}-\bar{X}\right)^{2}}{\sigma^{2}}
$$

there is $\mathbf{1}$ less degree of freedom, or a degree of freedom is lost in the estimation of $\mu$ (i.e., when $\mu$ is replaced by $\bar{x}$ ). In other words, there are $n$ degrees of freedom, or independent pieces of information, in the random sample from the normal distribution. When the data (the values in the sample) are used to compute the mean, there is 1 less degree of freedom in the information used to estimate $\sigma^{2}$.

## Theorem 8.4:

If $S^{2}$ is the variance of a random sample of size $n$ taken from a normal population having the variance $\sigma^{2}$, then the statistic

$$
\chi^{2}=\frac{(n-1) S^{2}}{\sigma^{2}}=\sum_{i=1}^{n} \frac{\left(X_{i}-\bar{X}\right)^{2}}{\sigma^{2}}
$$

has a chi-squared distribution with $v=n-1$ degrees of freedom.
$\chi_{\alpha}^{2}$ represent the $\chi^{2}$ value above which we find an area of $\alpha$.


Table A. 5 gives values of $\chi_{\alpha}^{2}$ for various values of $\alpha$ and $v$. The areas, $\alpha$, are the column headings; the degrees of freedom, $v$, are given in the left column; and the table entries are the $\chi^{2}$ values. Hence, the $\chi^{2}$ value with 7 degrees of freedom, leaving an area of 0.05 to the right, is $\chi_{0.05}^{2}=14.067$. Owing to lack of symmetry, we must also use the tables to find $\chi_{n 95}^{2}=2.167$ for $v=7$.

Table A. 5 Critical Values of the Chi-Squared Distribution


|  | $\alpha$ |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $v$ | 0.995 | 0.99 | 0.98 | 0.975 | 0.95 | 0.90 | 0.80 | 0.75 | 0.70 | 0.50 |
| 1 | $0.0^{4} 393$ | $0.0^{3} 157$ | $0.0^{3} 628$ | $0.0^{3} 982$ | 0.00393 | 0.0158 | 0.0642 | 0.102 | 0.148 | 0.455 |
| 2 | 0.0100 | 0.0201 | 0.0404 | 0.0506 | 0.103 | 0.211 | 0.446 | 0.575 | 0.713 | 1.386 |
| 3 | 0.0717 | 0.115 | 0.185 | 0.216 | 0.352 | 0.584 | 1.005 | 1.213 | 1.424 | 2.366 |
| 4 | 0.207 | 0.297 | 0.429 | 0.484 | 0.711 | 1.064 | 1.649 | 1.923 | 2.195 | 3.357 |
| 5 | 0.412 | 0.554 | 0.752 | 0.831 | 1.145 | 1.610 | 2.343 | 2.675 | 3.000 | 4.351 |
| 6 | 0.676 | 0.872 | 1.134 | 1.237 | 1.635 | 2.204 | 3.070 | 3.455 | 3.828 | 5.348 |
| 7 | 0.989 | 1.239 | 1.564 | 1.690 | 2.167 | 2.833 | 3.822 | 4.255 | 4.671 | 6.346 |
| 8 | 1.344 | 1.647 | 2.032 | 2.180 | 2.733 | 3.490 | 4.594 | 5.071 | 5.527 | 7.344 |
| 9 | 1.735 | 2.088 | 2.532 | 2.700 | 3.325 | 4.168 | 5.380 | 5.899 | 6.393 | 8.343 |
| 10 | 2.156 | 2.558 | 3.059 | 3.247 | 3.940 | 4.865 | 6.179 | 6.737 | 7.267 | 9.342 |
| 11 | 2.603 | 3.053 | 3.609 | 3.816 | 4.575 | 5.578 | 6.989 | 7.584 | 8.148 | 10.341 |
| 12 | 3.074 | 3.571 | 4.178 | 4.404 | 5.226 | 6.304 | 7.807 | 8.438 | 9.034 | 11.340 |
| 13 | 3.565 | 4.107 | 4.765 | 5.009 | 5.892 | 7.041 | 8.634 | 9.299 | 9.926 | 12.340 |
| 14 | 4.075 | 4.660 | 5.368 | 5.629 | 6.571 | 7.790 | 9.467 | 10.165 | 10.821 | 13.339 |
| 15 | 4.601 | 5.229 | 5.985 | 6.262 | 7.261 | 8.547 | 10.307 | 11.037 | 11.721 | 14.339 |
| 16 | 5.142 | 5.812 | 6.614 | 6.908 | 7.962 | 9.312 | 11.152 | 11.912 | 12.624 | 15.338 |
| 17 | 5.697 | 6.408 | 7.255 | 7.564 | 8.672 | 10.085 | 12.002 | 12.792 | 13.531 | 16.338 |
| 18 | 6.265 | 7.015 | 7.906 | 8.231 | 9.390 | 10.865 | 12.857 | 13.675 | 14.440 | 17.338 |
| 19 | 6.844 | 7.633 | 8.567 | 8.907 | 10.117 | 11.651 | 13.716 | 14.562 | 15.352 | 18.338 |
| 20 | 7.434 | 8.260 | 9.237 | 9.591 | 10.851 | 12.443 | 14.578 | 15.452 | 16.266 | 19.337 |
| 21 | 8.034 | 8.897 | 9.915 | 10.283 | 11.591 | 13.240 | 15.445 | 16.344 | 17.182 | 20.337 |
| 22 | 8.643 | 9.542 | 10.600 | 10.982 | 12.338 | 14.041 | 16.314 | 17.240 | 18.101 | 21.337 |
| 23 | 9.260 | 10.196 | 11.293 | 11.689 | 13.091 | 14.848 | 17.187 | 18.137 | 19.021 | 22.337 |
| 24 | 9.886 | 10.856 | 11.992 | 12.401 | 13.848 | 15.659 | 18.062 | 19.037 | 19.943 | 23.337 |
| 25 | 10.520 | 11.524 | 12.697 | 13.120 | 14.611 | 16.473 | 18.940 | 19.939 | 20.867 | 24.337 |
| 26 | 11.160 | 12.198 | 13.409 | 13.844 | 15.379 | 17.292 | 19.820 | 20.843 | 21.792 | 25.336 |
| 27 | 11.808 | 12.878 | 14.125 | 14.573 | 16.151 | 18.114 | 20.703 | 21.749 | 22.719 | 26.336 |
| 28 | 12.461 | 13.565 | 14.847 | 15.308 | 16.928 | 18.939 | 21.588 | 22.657 | 23.647 | 27.336 |
| 29 | 13.121 | 14.256 | 15.574 | 16.047 | 17.708 | 19.768 | 22.475 | 23.567 | 24.577 | 28.336 |
| 30 | 13.787 | 14.953 | 16.306 | 16.791 | 18.493 | 20.599 | 23.364 | 24.478 | 25.508 | 29.336 |
| 40 | 20.707 | 22.164 | 23.838 | 24.433 | 26.509 | 29.051 | 32.345 | 33.66 | 34.872 | 39.335 |
| 50 | 27.991 | 29.707 | 31.664 | 32.357 | 34.764 | 37.689 | 41.449 | 42.942 | 44.313 | 49.335 |
| 60 | 35.534 | 37.485 | 39.699 | 40.482 | 43.188 | 46.459 | 50.641 | 52.294 | 53.809 | 59.335 |


|  | $\alpha$ |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $v$ | 0.30 | 0.25 | 0.20 | 0.10 | 0.05 | 0.025 | 0.02 | 0.01 | 0.005 | 0.001 |
| 1 | 1.074 | 1.323 | 1.642 | 2.706 | 3.841 | 5.024 | 5.412 | 6.635 | 7.879 | 10.827 |
| 2 | 2.408 | 2.773 | 3.219 | 4.605 | 5.991 | 7.378 | 7.824 | 9.210 | 10.597 | 13.815 |
| 3 | 3.665 | 4.108 | 4.642 | 6.251 | 7.815 | 9.348 | 9.837 | 11.345 | 12.838 | 16.266 |
| 4 | 4.878 | 5.385 | 5.989 | 7.779 | 9.488 | 11.143 | 11.668 | 13.277 | 14.860 | 18.466 |
| 5 | 6.064 | 6.626 | 7.289 | 9.236 | 11.070 | 12.832 | 13.388 | 15.086 | 16.750 | 20.515 |
| 6 | 7.231 | 7.841 | 8.558 | 10.645 | 12.592 | 14.449 | 15.033 | 16.812 | 18.548 | 22.457 |
| 7 | 8.383 | 9.037 | 9.803 | 12.017 | 14.067 | 16.013 | 16.622 | 18.475 | 20.278 | 24.321 |
| 8 | 9.524 | 10.219 | 11.030 | 13.362 | 15.507 | 17.535 | 18.168 | 20.090 | 21.955 | 26.124 |
| 9 | 10.656 | 11.389 | 12.242 | 14.684 | 16.919 | 19.023 | 19.679 | 21.666 | 23.589 | 27.877 |
| 10 | 11.781 | 12.549 | 13.442 | 15.987 | 18.307 | 20.483 | 21.161 | 23.209 | 25.188 | 29.588 |
| 11 | 12.899 | 13.701 | 14.631 | 17.275 | 19.675 | 21.920 | 22.618 | 24.725 | 26.757 | 31.264 |
| 12 | 14.011 | 14.845 | 15.812 | 18.549 | 21.026 | 23.337 | 24.054 | 26.217 | 28.300 | 32.909 |
| 13 | 15.119 | 15.984 | 16.985 | 19.812 | 22.362 | 24.736 | 25.471 | 27.688 | 29.819 | 34.527 |
| 14 | 16.222 | 17.117 | 18.151 | 21.064 | 23.685 | 26.119 | 26.873 | 29.141 | 31.319 | 36.124 |
| 15 | 17.322 | 18.245 | 19.311 | 22.307 | 24.996 | 27.488 | 28.259 | 30.578 | 32.801 | 37.698 |
| 16 | 18.418 | 19.369 | 20.465 | 23.542 | 26.296 | 28.845 | 29.633 | 32.000 | 34.267 | 39.252 |
| 17 | 19.511 | 20.489 | 21.615 | 24.769 | 27.587 | 30.191 | 30.995 | 33.409 | 35.718 | 40.791 |
| 18 | 20.601 | 21.605 | 22.760 | 25.989 | 28.869 | 31.526 | 32.346 | 34.805 | 37.156 | 42.312 |
| 19 | 21.689 | 22.718 | 23.900 | 27.204 | 30.144 | 32.852 | 33.687 | 36.191 | 38.582 | 43.819 |
| 20 | 22.775 | 23.828 | 25.038 | 28.412 | 31.410 | 34.170 | 35.020 | 37.566 | 39.997 | 45.314 |
| 21 | 23.858 | 24.935 | 26.171 | 29.615 | 32.671 | 35.479 | 36.343 | 38.932 | 41.401 | 46.796 |
| 22 | 24.939 | 26.039 | 27.301 | 30.813 | 33.924 | 36.781 | 37.659 | 40.289 | 42.796 | 48.268 |
| 23 | 26.018 | 27.141 | 28.429 | 32.007 | 35.172 | 38.076 | 38.968 | 41.638 | 44.181 | 49.728 |
| 24 | 27.096 | 28.241 | 29.553 | 33.196 | 36.415 | 39.364 | 40.270 | 42.980 | 45.558 | 51.179 |
| 25 | 28.172 | 29.339 | 30.675 | 34.382 | 37.652 | 40.646 | 41.566 | 44.314 | 46.928 | 52.619 |
| 26 | 29.246 | 30.435 | 31.795 | 35.563 | 38.885 | 41.923 | 42.856 | 45.642 | 48.290 | 54.051 |
| 27 | 30.319 | 31.528 | 32.912 | 36.741 | 40.113 | 43.195 | 44.140 | 46.963 | 49.645 | 55.475 |
| 28 | 31.391 | 32.620 | 34.027 | 37.916 | 41.337 | 44.461 | 45.419 | 48.278 | 50.994 | 56.892 |
| 29 | 32.461 | 33.711 | 35.139 | 39.087 | 42.557 | 45.722 | 46.693 | 49.588 | 52.335 | 58.301 |
| 30 | 33.530 | 34.800 | 36.250 | 40.256 | 43.773 | 46.979 | 47.962 | 50.892 | 53.672 | 59.702 |
| 40 | 44.165 | 45.616 | 47.269 | 51.805 | 55.758 | 59.342 | 60.436 | 63.691 | 66.766 | 73.403 |
| 50 | 54.723 | 56.334 | 58.164 | 63.167 | 67.505 | 71.420 | 72.613 | 76.154 | 79.490 | 86.660 |
| 60 | 65.226 | 66.981 | 68.972 | 74.397 | 79.082 | 83.298 | 84.58 | 88.379 | 91.952 | 99.608 |

Exactly $95 \%$ of a chi-squared distribution lies between $\chi_{0.975}^{2}$ and $\chi_{0.025}^{2}$. A $\chi^{2}$ value falling to the right of $\chi_{0.025}^{2}$ is not likely to occur unless our assumed value of $\sigma^{2}$ is too small. Similarly, a $\chi^{2}$ value falling to the left of $\chi_{0.975}^{2}$ is unlikely unless our assumed value of $\sigma^{2}$ is too large. In other words, it is possible to have a $\chi^{2}$ value to the left of $\chi_{0.975}^{2}$ or to the right of $\chi_{0.025}^{2}$ when $\sigma^{2}$ is correct, but if this should occur, it is more probable that the assumed value of $\sigma^{2}$ is in error.

## Example 8.7:

A manufacturer of car batteries guarantees that the batteries will last, on average, 3 years with a standard deviation of 1 year. If five of these batteries have lifetimes of $1.9,2.4,3.0,3.5$, and 4.2 years, should the manufacturer still be convinced that the batteries have a standard deviation of 1 year? Assume that the battery lifetime follows a normal distribution.

## Solution :

We first find the sample variance using Theorem 8.1,

$$
s^{2}=\frac{(5)(48.26)-(15)^{2}}{(5)(4)}=0.815 .
$$

Then

$$
\chi^{2}=\frac{(4)(0.815)}{1}=3.26
$$

is a value from a chi-squared distribution with 4 degrees of freedom. Since $95 \%$ of the $\chi^{2}$ values with 4 degrees of freedom fall between 0.484 and 11.143 , the computed value with $\sigma^{2}=1$ is reasonable, and therefore the manufacturer has no reason to suspect that the standard deviation is other than 1 year.


